Introductory Econometrics

Simple Linear Regression Model (II)

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Recall: Bivariate Linear Regression Model

• Recall the bivariate linear regression model

$$Y = \beta_0 + \beta_1 X + u,$$

where β_0 and β_1 are parameters to be estimated, and

- β_0 is referred to as the **intercept**.
- β_1 is referred to as the **slope**.
- We observe Y and X as **sample**,

 $\{(X_i, Y_i) : i = 1, 2, \dots, n\}$

• For each *i*, bivariate linear regression model suggests

$$Y_i = \beta_0 + \beta_1 X_i + u_i.$$

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Normal equations

$$\begin{cases} \sum \left(Y_i - \hat{\beta}_0 - \hat{\beta}_1 X_i \right) = 0\\ \sum X_i \left(Y_i - \hat{\beta}_0 - \hat{\beta}_1 X_i \right) = 0 \end{cases}$$

Solution

$$\begin{cases} \hat{\beta}_0 = \frac{\sum X_i^2 \sum Y_i - \sum X_i \sum X_i Y_i}{n \sum X_i^2 - (\sum X_i)^2} \\ \hat{\beta}_1 = \frac{n \sum Y_i X_i - \sum Y_i \sum X_i}{n \sum X_i^2 - (\sum X_i)^2} \end{cases}$$

Assumptions for Simple Linear Regression

Assumption 1: Model is correctly specified.

- (1) Variables are correctly selected.
- (2) Model is correctly specified.

Assumption 2: Explanatory variable converges in probability to constant.

$$P_{n \to \infty} \sum_{i=1}^{n} \left(X_i - \bar{X} \right)^2 / n \to Q.$$

Assumption 3: Stochastic error has zero expectation.

$$\mathrm{E}\left(u_{i} \mid X\right) = 0.$$

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 $E(u_i | X) = 0$ suggests that $Cov(u_i, X) = 0$, i.e. u_i is uncorrelated with X. We need an important tool called **Law of Iterated Expectations, LIE**.

Assumption 4: Homeskedasticity and no correlation.

$$Var(u_i | X) = \sigma^2 \quad i = 1, 2, ..., n$$

 $\operatorname{Cov}\left(u_{i}, u_{j} \mid X\right) = 0 \quad i \neq j$

And similarly due to LIE, we have

$$\operatorname{Var}(u_i) = \sigma^2.$$

Assumptions for Simple Linear Regression

Assumption 5: Normality.

 $u_i \mid X \sim N\left(0, \sigma^2\right).$

Assumption 1 to Assumption 5 are usually referred to as the classical assumption, and the corresponding linear regression model is referred as to the Classical Linear Regression Model, CLRM.

- Specifically, Assumption 5 can be relaxed in large sample, Assumption 1 to Assumption 4 refers to the Gauss-Markov assumption.
- For bivariate linear regression model

$$Y \mid X \sim N\left(\beta_0 + \beta_1 X, \sigma^2\right).$$

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• Statistics and statistical properties of statistics.

Small-sample-properties:

- Unbiasedness
- Efficiency

Large-sample-properties:

- Consistency
- $\hat{\theta}$ is unbiased if $E(\hat{\theta}) = \theta$, that is the average value of $\hat{\theta}$ over all realizations is equal to the underlying population value.
- $\hat{\theta}$ is efficient if it has smaller variance than **any other unbiased estimator**.

Convergence in Probability. A sequence of random variables {X_n} converges in probability to a random variable X if, for every ε > 0,

$$\lim_{n \to \infty} \mathbb{P}\left(\|X_n - X\| \ge \epsilon \right) = 0.$$

Notation: $X_n \xrightarrow{p} X$, or $P \lim (X_n) = X$.

• An estimator $\hat{\theta}$ is consistent if $P \lim (\hat{\theta}) = \theta$.

• OLS estimator is a linear combination of Y_i .

$$\hat{\beta}_1 = \sum k_i Y_i \quad \hat{\beta}_0 = \sum w_i Y_i$$

where $k_i = \frac{x_i}{\sum x_i^2}$ and $w_i = \frac{1}{n} - \bar{X}k_i$.

- OLS estimator is unbiased.
- OLS estimator is efficient. More discussion
- OLS estimator is **best linear unbiased estimator (BLUE)**. This claim is usually referred to as the **Gauss-Markov theorem**.

Large-Sample Properties of OLS Estimator

- Large sample properties refer to the properties of statistics when sample size is "sufficiently large", i.e. $n \to \infty$.
- Weak Law of Large Numbers. Suppose X_1, \ldots, X_n and *i.i.d.* with $E ||X_1|| < \infty$, then as $n \to \infty$,

$$\bar{X}_n = \frac{1}{n} \sum_{i=1}^n X_i \xrightarrow{p} \mathcal{E}(X_1).$$

• OLS estimator is consistent.

$$P \lim \left(\hat{\beta}_{1}\right) = P \lim \left(\beta_{1} + \sum k_{i}u_{i}\right) = P \lim \left(\beta_{1}\right) + P \lim \left(\frac{\sum x_{i}u_{i}}{\sum x_{i}^{2}}\right)$$
$$= \beta_{1} + P \lim \left(\frac{\sum x_{i}u_{i}/n}{\sum x_{i}^{2}/n}\right) = \beta_{1}.$$

• If we stick to the Normality assumption (Assumption 5), it can be easily shown that for the finite sample

$$\hat{\beta}_1 \sim N\left(\beta_1, \frac{\sigma^2}{\sum x_i^2}\right)$$
$$\hat{\beta}_0 \sim N\left(\beta_0, \frac{\sum X_i^2}{n \sum x_i^2} \sigma^2\right)$$

• By relaxing the normality assumption under the large sample, assymptotically we have

$$\hat{\beta}_1 \stackrel{a}{\sim} N\left(\beta_1, \frac{\sigma^2}{\sum x_i^2}\right)$$
$$\hat{\beta}_0 \stackrel{a}{\sim} N\left(\beta_0, \frac{\sum X_i^2}{n \sum x_i^2} \sigma^2\right)$$

Estimation of σ^2

$$\sum e_i^2 = \sum (y_i - \hat{y}_i)^2$$

=
$$\sum [(\beta_1 - \hat{\beta}_1) x_i + (u_i - \bar{u})]^2$$

=
$$\sum (\beta_1 - \hat{\beta}_1)^2 x_i^2 + \sum (u_i - \bar{u})^2 - 2 \sum (\sum k_i u_i) x_i (u_i - \bar{u})$$

• Note that

$$-2\sum \left(\sum k_{i}u_{i}\right)x_{i}\left(u_{i}-\bar{u}\right) = -2\sum x_{i}u_{i}\sum k_{i}u_{i}+2\bar{u}\sum x_{i}\sum k_{i}u_{i}$$
$$= -2\sum x_{i}u_{i}\sum k_{i}u_{i}$$

• Hence

$$\sum e_i^2 = \sum \left(\beta_1 - \hat{\beta}_1\right)^2 x_i^2 + \sum \left(u_i - \bar{u}\right)^2 - 2 \sum x_i u_i \frac{\sum x_i u_i}{\sum x_i^2}$$

Estimation of σ^2

$$\mathbb{E}\left[\sum \left(\beta_1 - \hat{\beta}_1\right)^2 x_i^2 \mid X\right] = \sum x_i^2 \operatorname{Var}\left(\hat{\beta}_1 \mid X\right) = \frac{\sum x_i^2 \sigma^2}{\sum x_i^2} = \sigma^2$$

$$E\left[\sum (u_i - \bar{u})^2 \mid X\right] = E\left[\left(\sum u_i^2 - 2\bar{u}\sum u_i + n\bar{u}^2\right) \mid X\right]$$
$$= E\left[\left(\sum u_i^2 - n\bar{u}^2\right) \mid X\right] = (n - 1)\sigma^2$$

$$\mathbf{E}\left[\left.\frac{\left(\sum x_{i}u_{i}\right)^{2}}{\sum x_{i}^{2}}\right|X\right] = \mathbf{E}\left[\left.\frac{\sum x_{i}^{2}u_{i}^{2} + 2\sum_{i\neq j}\left(x_{i}x_{j}\right)\left(u_{i}u_{j}\right)}{\sum x_{i}^{2}}\right|X\right] = \sigma^{2}$$

• Finally

$$E\left(\sum e_i^2 \mid X\right) = \sigma^2 + (n-1)\sigma^2 - 2\sigma^2 = (n-2)\sigma^2.$$

• We denote the estimation of σ^2 by $\hat{\sigma}^2$ as follows

$$\hat{\sigma}^2 = \frac{\sum e_i^2}{n-2}.$$

• With $\hat{\sigma}^2$, we obtain the sample variance of $\hat{\beta}_1$ and $\hat{\beta}_0$ respectively as follows

$$S_{\hat{\beta}_{1}}^{2} = \hat{\sigma}^{2} / \sum x_{i}^{2}$$
$$S_{\hat{\beta}_{0}}^{2} = \hat{\sigma}^{2} \sum X_{i}^{2} / n \sum x_{i}^{2}$$

- Classical assumptions for simple linear regression.
- Law of Iterated Expectations.
- Statistical properties of OLS estimator.
- Corresponding Distributions.

Variance of $\hat{\beta}_1$

$$\operatorname{Var}\left(\hat{\beta}_{1} \mid X\right) = \operatorname{Var}\left(\sum k_{i}Y_{i} \mid X\right)$$
$$= \sum k_{i}^{2}\operatorname{Var}\left[\left(\beta_{0} + \beta_{1}X_{i} + u_{i}\right) \mid X\right]$$
$$= \sum k_{i}^{2}\operatorname{Var}\left(u_{i} \mid X\right)$$
$$= \sum \left(\frac{x_{i}}{\sum x_{i}^{2}}\right)^{2}\sigma^{2} = \frac{\sigma^{2}}{\sum x_{i}^{2}}$$

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